

Effect of Torsemide vs Furosemide After Discharge on All-Cause Mortality in Patients Hospitalized With Heart Failure

The TRANSFORM-HF Randomized Clinical Trial

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IMPORTANCE Although furosemide is the most commonly used loop diuretic in patients with heart failure, some studies suggest a potential benefit for torsemide.

OBJECTIVE To determine whether torsemide results in decreased mortality compared with furosemide among patients hospitalized for heart failure.

DESIGN, SETTING, AND PARTICIPANTS TRANSFORM-HF was an open-label, pragmatic randomized trial that recruited 2859 participants hospitalized with heart failure (regardless of ejection fraction) at 60 hospitals in the United States. Recruitment occurred from June 2018 through March 2022, with follow-up through 30 months for death and 12 months for hospitalizations. The final date for follow-up data collection was July 2022.

INTERVENTIONS Loop diuretic strategy of torsemide (n = 1431) or furosemide (n = 1428) with investigator-selected dosage.

MAIN OUTCOMES AND MEASURES The primary outcome was all-cause mortality in a time-to-event analysis. There were 5 secondary outcomes with all-cause mortality or all-cause hospitalization and total hospitalizations assessed over 12 months being highest in the hierarchy. The prespecified primary hypothesis was that torsemide would reduce all-cause mortality by 20% compared with furosemide.

RESULTS TRANSFORM-HF randomized 2859 participants with a median age of 65 years (IQR, 56-75), 36.9% were women, and 33.9% were Black. Over a median follow-up of 17.4 months, a total of 113 patients (53 [3.7%] in the torsemide group and 60 [4.2%] in the furosemide group) withdrew consent from the trial prior to completion. Death occurred in 373 of 1431 patients (26.1%) in the torsemide group and 374 of 1428 patients (26.2%) in the furosemide group (hazard ratio, 1.02 [95% CI, 0.89-1.18]). Over 12 months following randomization, all-cause mortality or all-cause hospitalization occurred in 677 patients (47.3%) in the torsemide group and 704 patients (49.3%) in the furosemide group (hazard ratio, 0.92 [95% CI, 0.83-1.02]). There were 940 total hospitalizations among 536 participants in the torsemide group and 987 total hospitalizations among 577 participants in the furosemide group (rate ratio, 0.94 [95% CI, 0.84-1.07]). Results were similar across prespecified subgroups, including among patients with reduced, mildly reduced, or preserved ejection fraction.

CONCLUSIONS AND RELEVANCE Among patients discharged after hospitalization for heart failure, torsemide compared with furosemide did not result in a significant difference in all-cause mortality over 12 months. However, interpretation of these findings is limited by loss to follow-up and participant crossover and nonadherence.

TRIAL REGISTRATION ClinicalTrials.gov Identifier: [NCT03296813](https://clinicaltrials.gov/ct2/show/study/NCT03296813)

JAMA. 2023;329(3):214-223. doi:10.1001/jama.2022.23924

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Heat failure is a major and growing public health problem worldwide.¹ Many patients with heart failure experience symptoms of congestion and volume overload including dyspnea and edema.² The majority of patients with symptomatic heart failure are prescribed loop diuretics for the treatment of congestion.³⁻⁶ Guidelines indicate that the use of diuretics is a cornerstone of a successful approach to the treatment of congestion in heart failure.⁷

Furosemide is the most commonly used loop diuretic for heart failure.^{5,8} However, preclinical and clinical data suggest potential benefits of torsemide compared with furosemide. Torsemide has increased bioavailability and a longer half-life than furosemide.³ Torsemide may also have beneficial effects on myocardial fibrosis, aldosterone production, sympathetic activation, ventricular remodeling, and natriuretic peptides.^{9,10} Several small studies of torsemide vs furosemide and meta-analyses suggest a decrease in morbidity and potentially mortality with torsemide compared with furosemide.^{8,11-13} However, in light of the lack of an adequately powered clinical outcomes study, there is insufficient evidence to recommend torsemide over furosemide.

The TRANSFORM-HF (Torsemide Comparison With Furosemide for Management of Heart Failure) Trial was designed to compare the effect of torsemide with furosemide in patients hospitalized with heart failure. The trial was an open-label, pragmatic, randomized, comparative-effectiveness study to assess whether a strategy of torsemide vs furosemide on discharge from the hospital would result in a lower risk of death from any cause among patients with heart failure (regardless of ejection fraction).¹⁴ The study incorporated pragmatic elements to perform a real-world trial of loop diuretic strategies in routine heart failure care.¹⁵ The primary hypothesis was that torsemide would reduce all-cause mortality by 20% compared with furosemide.

Methods

Trial Design and Oversight

The trial design and operations have been previously described.¹⁴ In brief, the event-driven trial was conducted in 60 hospitals in the United States. Patients were recruited during hospitalization with heart failure. The trial protocol and statistical analysis plan are available in [Supplement 1](#) and [Supplement 2](#), respectively. The protocol was designed by academic investigators at Duke Clinical Research Institute, Yale School of Medicine, and the University of Michigan in collaboration with the National Heart, Lung, and Blood Institute. An independent data and safety monitoring board (DSMB) approved the trial protocol and monitored patient safety throughout the trial. The trial was approved by the Duke University institutional review board (IRB) as well as a central IRB or local site IRBs. All patients provided written informed consent before enrollment.

Trial Participants

Eligible patients were hospitalized for heart failure and could have either de novo heart failure or worsening of chronic heart

Key Points

Question Does torsemide reduce all-cause mortality compared with furosemide in patients with heart failure following hospitalization?

Findings In this randomized clinical trial of 2859 patients, 26.1% of patients randomized to torsemide and 26.2% randomized to furosemide died over a median follow-up of 17.4 months without a significant difference between groups.

Meaning Among patients discharged after hospitalization for heart failure, torsemide compared with furosemide did not result in a significant difference in all-cause mortality over 12 months; however, interpretation of these findings is limited by loss to follow-up and participant crossover and nonadherence.

failure as defined by the inclusion and exclusion criteria in the eMethods in [Supplement 3](#). In brief, participants had either a left ventricular ejection fraction of 40% or less within 24 months or an elevated natriuretic peptide level during the index hospitalization as measured by the local laboratory. Participants needed to have a plan for daily outpatient oral loop diuretic with anticipated long-term use. Patients were recruited during hospitalization up until the time of discharge. Patients with end-stage kidney disease requiring dialysis or a history of heart transplant or left ventricular assist device were excluded. The inclusion of race and ethnicity data was aligned with National Institutes of Health guidance. Participants made the determination based on fixed categories, which allowed multiple responses and “other.”

Randomization

Treatment assignment was generated using a simple randomization scheme (ie, no stratification) given the open-label nature of the intervention to limit the potential bias due to predictable treatment assignment.

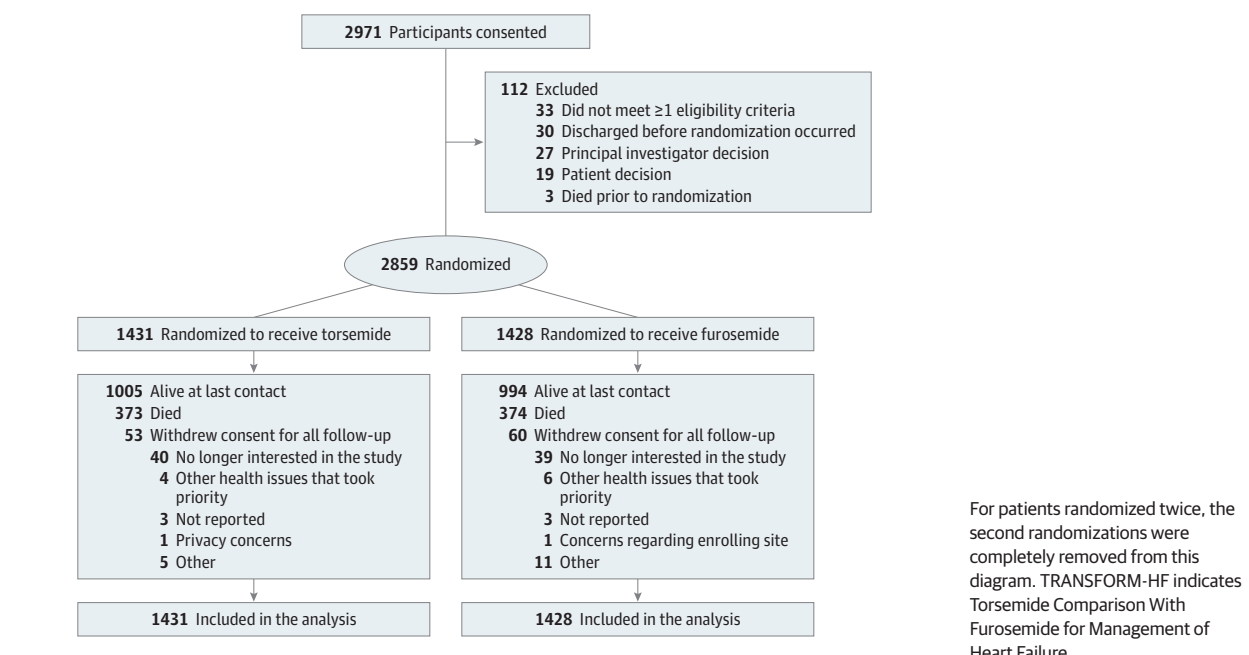
Interventions

After providing informed consent, patients were randomly assigned in a 1:1 ratio to a treatment strategy of torsemide or furosemide prior to hospital discharge ([Figure 1](#)). Dose and frequency of the randomized therapy during hospitalization and at discharge were determined by the treating clinician with the following conversion provided as a guide: 1 mg of torsemide to 2 to 4 mg of oral furosemide. Flexibility in dosing 2 to 4 mg of furosemide vs 1 mg of torsemide was chosen given the low quality and somewhat contradictory data available regarding dose conversion of these diuretics. Changes in dose and frequency of the randomized therapy after discharge were at the discretion of the patient’s usual outpatient clinicians. Participants received medication as per routine care with open-label prescription. Participants were provided medication adherence and educational materials to support adherence to the randomized therapy.

Follow-up

After this initial contact, no further study-specific patient contact was required at the site level. The trial used centralized

Figure 1. Participant Flow in the TRANSFORM-HF Randomized Clinical Trial



follow-up via the Duke Clinical Research Institute call center. Participants had telephone interviews at 30 days, 6 months, and 12 months following discharge. To provide estimates on long-term treatment adherence and to support adequate event accrual, the first 1500 participants had additional follow-up: the first 500 participants received telephone calls every 6 months through 30 months, participants 501 through 1000 received calls through 24 months, and participants 1001 through 1500 received telephone calls through 18 months.

At each telephone interview, participants (or approved proxies) provided information on adherence to the randomized therapy, vital status, hospitalization events, and patient-reported measures. Hospitalization events were verified by the call center using hospitalization records when possible. The National Death Index (NDI) was searched at regular intervals to confirm deaths and supplement vital status data obtained by the call center as previously detailed.¹⁶

The trial was event driven and designed to continue until at least 721 primary end points events (all-cause mortality) were observed with an initially projected enrollment of approximately 6000 patients. Interim data analyses for efficacy were conducted by the DSMB due to the large sample size. The DSMB met approximately every 6 months to review study progress.

Trial Outcomes

The primary effectiveness outcome was all-cause mortality assessed in a time-to-event analysis. All-cause mortality was selected based on prior data supporting a reduction in death with torsemide and the need to minimize bias in the setting of an unblinded trial. The prespecified subgroups were age (<65 , ≥ 65 ; <75 , ≥ 75 years), sex, race and ethnicity (Asian, Black, White, other), ejection fraction ($\leq 40\%$, 41% - 49% , $\geq 50\%$), loop diuretic treatment prior to index hospitalization, New York

Heart Association class at randomization (I/II vs III/IV), systolic blood pressure ($</\geq$ median), estimated glomerular filtration rate categories (<30 , ≥ 30 to <60 , ≥ 60 mL/min/1.73 m²), diabetes (yes/no), mineralocorticoid receptor antagonist use at randomization (yes/no), academic/university hospital (yes/no), and duration of heart failure. There were 5 secondary end points, with 3 of them being clinical outcomes with a hierarchy of all-cause mortality or all-cause hospitalization over 12 months, total hospitalizations assessed over 12 months, and all-cause mortality or all-cause hospitalization assessed over 30 days. All-cause hospitalization was evaluated as opposed to cardiovascular or heart failure hospitalization to assess the total readmission burden. The 2 remaining secondary outcomes were quality-of-life end points at 12 months and will be reported separately.

Data Sources

Outcomes were ascertained from multiple data sources, including patient (or proxy) report at scheduled trial encounters, queries of medical records, obituaries, grave markers, and the NDI. NDI assessments were performed through December 2021. Details on the process for triggering and verifying outcomes and censoring rules are described in [Supplement 2](#). In brief, the NDI is the most complete death data set available in the United States.¹⁶ The implication is that NDI deaths are actual deaths and the absence of an NDI death means that a patient may be considered alive at the end of the reporting year. Therefore, use of the NDI supported very low rates of missing data for the primary end point.

Sample Size Calculation

The trial protocol and statistical analysis plan in [Supplement 1](#) and [Supplement 2](#), respectively, provide details on the sample

size calculation. In brief, the planned study population for this trial was broader than most prior heart failure clinical trials. As such, it was difficult to anticipate the expected event rates. However, prior observational data demonstrated 1-year mortality rates following heart failure hospitalization of more than 30%.¹⁷ A meta-analysis of studies assessing mortality with torsemide vs furosemide demonstrated a nominal reduction in death events of more than 20% with torsemide.⁸ Therefore, the required number of primary end point events (ie, all-cause mortality) necessary to obtain power ranging from 80% to 90% with hazard ratios ranging from 0.75 to 0.85 was evaluated (Supplement 2). At least 721 primary end point events were needed to have 85% power or more to detect a hazard ratio of 0.80 assuming 1:1 randomization, a 2-side type I error of .05, and a test statistic based on the log-rank test. It was initially estimated that up to 6000 participants would be needed to accrue the necessary event count. Power calculations did not account for inclusion of covariates in the primary outcome model.

Statistical Analysis

The full analysis set included all randomized patients and was the primary analysis population. Comparisons based on randomized treatment assignments were performed. Descriptive summaries of baseline variables were generated for each randomized treatment group. Continuous variables were presented as medians with IQRs or means with the SDs, and discrete variables were summarized with the use of frequencies and percentages.

For the primary end point of all-cause mortality, a Cox proportional hazards regression model was used to assess outcome differences between the 2 treatment groups and compute a hazard ratio and 95% CI. Prespecified covariates in the primary model included randomized treatment, age, sex, ejection fraction category ($\leq 40\%$, 41%-49%, $\geq 50\%$), and loop diuretic treatment before index hospital admission. Additional post hoc analyses were performed to (1) include site as a random effect in the adjusted model, (2) report the unadjusted Cox model data with robust variance estimators, and (3) assess the implications of the COVID-19 pandemic via a time-dependent covariate before or after the US national emergency data of March 13, 2020. All events post randomization (including in-hospital deaths) were included in the analysis. The proportional hazards assumption was assessed using the ZPH test option in the SAS Cox PH regression model statement (PROC PHREG), which is based on weighted Schoenfeld residuals. There was no violation in proportional hazards. As detailed in the statistical analysis plan (section 12.1, Supplement 2), censoring based on the NDI follow-up period was the primary censoring definition. Regarding the baseline covariates in the outcome model, no patients were excluded due to missing data. Age and sex were available for all patients and for ejection fraction and prior loop diuretic categories, an “unknown category” was used so there were no missing data.

With regard to secondary end points, analyses of the composite of all-cause mortality or all-cause hospitalization at 30 days and 12 months was by time-to-event as for the primary end point analysis. The frequency of primary all-cause rehospitalization events was analyzed by negative binomial

regression with relative risks and 95% CIs provided. To address competing risk, a post hoc analysis was performed with multivariable Fine and Gray competing risk models for all-cause hospitalizations through 12 months in the full analysis data set as well as on-treatment at discharge and 30 days.

Prespecified supportive analyses were based on the subset of participants discharged alive taking the assigned medication (as-treated at discharge) and as-treated at day 30 as detailed in Supplement 2.

All analyses were performed with the use of SAS software version 9.4 (SAS Institute). For the primary analysis, a 2-sided *P* value less than .05 was considered statistically significant. For all other analyses, including secondary analyses and subgroup analyses, a *P* value less than .005 was considered statistically significant to improve the reproducibility of study results.¹⁸ *P* values are only reported until the last comparison for which the *P* value is significant. Thus, *P* values for the first nonsignificant comparison and for all comparisons thereafter are not reported. The secondary end points in section 4.2 of the statistical analysis plan (Supplement 2) are listed in the order of importance and testing. The widths of the 95% CIs are not adjusted for multiplicity and the intervals should not be used in place of hypothesis testing. The interpretation of these confidence intervals avoids the language of definitive conclusions used to report statistically significant findings as assessed by formal hypothesis testing.

Results

Patients and Follow-up

Recruitment began in June 2018. Following a routine DSMB meeting on February 18, 2022, the DSMB recommended stopping recruitment because the sample size was sufficient to answer the primary research question. The trial sponsor (National Heart, Lung, and Blood Institute) reviewed and accepted these recommendations with determination that the trial should execute an orderly closeout.

Recruitment ended March 4, 2022, with 2859 randomized participants (1431 to torsemide and 1428 to furosemide). Following the recommendation of the DSMB and sponsor to conclude the trial, participants recruited in the past 12 months received an end-of-study contact from the call center by May 15, 2022. The final date for follow-up data collection was July 29, 2022.

The characteristics of the patients at baseline were similar in the 2 groups (Table 1). The median age of participants was 65 years (IQR, 56-75); 36.9% of participants were women and 33.9% were Black. In the subset with reduced ejection fraction heart failure ($\leq 40\%$, $n = 1836$), baseline β -blocker use was 81.5%, angiotensin-converting enzyme/angiotensin receptor blocker or angiotensin receptor-neprilysin inhibitor (ARNI) use was 67.5% (25.2% ARNI use), mineralocorticoid receptor antagonist use was 44.3%, and sodium-glucose cotransporter 2 (SGLT2) inhibitor use was 7.8%. A total of 113 patients (53 [3.7%] in the torsemide group and 60 [4.2%] in the furosemide group) withdrew consent from the trial prior to completion (Figure 1). The median duration of follow-up

Table 1. Baseline Characteristics of TRANSFORM-HF Participants by Treatment Group

Characteristic	No. (%) ^a	
	Torsemide (n = 1431)	Furosemide (n = 1428)
Age, y		
Mean (SD)	64.0 (14.0)	65.0 (14.0)
Median (IQR)	65.0 (55.0-74.0)	65.5 (56.0-75.0)
Sex		
Female	498 (34.8)	557 (39.0)
Male	933 (65.2)	871 (61.0)
Race ^b		
American Indian or Alaska Native	9 (0.6)	3 (0.2)
Asian	37 (2.6)	26 (1.8)
Black or African American	474 (33.1)	494 (34.6)
Native Hawaiian or Pacific Islander	13 (0.9)	7 (0.5)
White	831 (58.1)	837 (58.6)
Other	44 (3.1)	35 (2.5)
Multiple	21 (1.5)	23 (1.6)
Not reported	2 (0.1)	3 (0.2)
Hispanic ethnicity, No./total (%)	75/1430 (5.2)	80/1425 (5.6)
Newly diagnosed heart failure	428 (29.9)	410 (28.7)
Heart failure hospitalization in past year, No./total (%)	524/1415 (37.0)	476/1414 (33.7)
Left ventricular ejection fraction, No./total (%), %		
≥50	318/1334 (23.8)	330/1301 (25.4)
41-49	81/1334 (6.1)	70/1301 (5.4)
≤40	935/1334 (70.1)	901/1301 (69.3)
Prior loop diuretic (before randomization)		
Furosemide	754 (52.7)	778 (54.5)
Torsemide	146 (10.2)	113 (7.9)
Bumetanide	64 (4.5)	65 (4.6)
Ischemic etiology	427 (29.8)	381 (26.7)
Comorbidities		
Atrial fibrillation/flutter	625/1419 (44.0)	649/1420 (45.7)
Diabetes	688 (48.1)	676 (47.3)
Chronic kidney disease	497 (34.7)	512 (35.9)
Vital signs		
Systolic blood pressure, mm Hg	118 (19)	119 (20)
Heart rate, /min	81 (16)	80 (16)
No.	1430	1427
Body mass index ^c	32.3 (9.7)	32.0 (9.3)
Baseline laboratories, median (IQR)		
NT-proBNP, pg/mL	3994 (1938-8850)	3833 (1936-7807)
No.	680	696
BNP, pg/dL	982 (468-1790)	921 (480-1865)
No.	703	678
Estimated GFR, mean (SD), mL/min/1.73 m ²	59.1 (25.0)	59.7 (26.0)
No.	1429	1425
Devices and medications		
β-Blocker	1140 (79.7)	1106 (77.5)
ACE inhibitor or ARB	640 (44.7)	603 (42.2)
Mineralocorticoid receptor antagonist	524 (36.6)	498 (34.9)
Sacubitril-valsartan	264 (18.4)	272 (19.0)
SGLT2 inhibitor	89/1383 (6.4)	81/1375 (5.9)
Implantable cardioverter-defibrillator	293/1428 (20.5)	298/1426 (20.9)
Cardiac resynchronization therapy	119/1430 (8.3)	105/1427 (7.4)

Abbreviations: ACE, angiotensin-converting enzyme; ARB, angiotensin receptor blocker; BNP, B-type natriuretic peptide; GFR, glomerular filtration rate; NT-proBNP, N-terminal pro-brain natriuretic peptide; SGLT2, sodium glucose cotransporter 2.

^a Values shown as No. (%) or mean (SD), unless otherwise specified.

^b The inclusion of race and ethnicity data was aligned with National Institutes of Health guidance. Participants made the determination based on fixed categories, which allowed multiple responses and "other."

^c Calculated as weight in kilograms divided by height in meters squared.

for all-cause mortality was 17.4 months (IQR, 8.0-29.0) and was similar in the 2 groups.

Primary Outcome

Death occurred in 373 of 1431 patients (26.1%) in the toremide group and 374 of 1428 patients (26.2%) in the furosemide group (hazard ratio, 1.02 [95% CI, 0.89-1.18]; $P = .76$) (Figure 2 and Table 2). There were 11 deaths during index hospitalization (7 in the toremide group and 4 in the furosemide group). The effect of toremide on the primary outcome was consistent across prespecified subgroups (Figure 3). Prespecified sensitivity analyses in the as-treated population were consistent with the primary analysis (eTable 1 in Supplement 3) and baseline characteristics by adherence are reported in eTable 2 in Supplement 3. Post hoc analyses that included site as a random effect in the adjusted model, the unadjusted Cox model data with robust variance estimators, and a COVID-19 assessment were consistent with the primary results (eTable 3 and eTable 4 in Supplement 3).

Secondary Outcomes

All-cause mortality or all-cause hospitalization occurred in 677 patients (47.3%) in the toremide group and 704 patients (49.3%) in the furosemide group (hazard ratio, 0.92 [95% CI, 0.83-1.02]) (eFigure in Supplement 3; Table 2). There were 940 total hospitalizations among 536 participants (37.5%) in the toremide group and 987 total hospitalizations among 577 participants (40.4%) in the furosemide group (rate ratio, 0.94 [95% CI, 0.84-1.07]) (Table 2). The post hoc analysis to address competing risk for all-cause hospitalizations through 12 months demonstrated a hazard ratio of 0.88 (95% CI, 0.78-0.99) with toremide compared with furosemide with consistent results in the on-treatment groups (eTable 5 in Supplement 3).

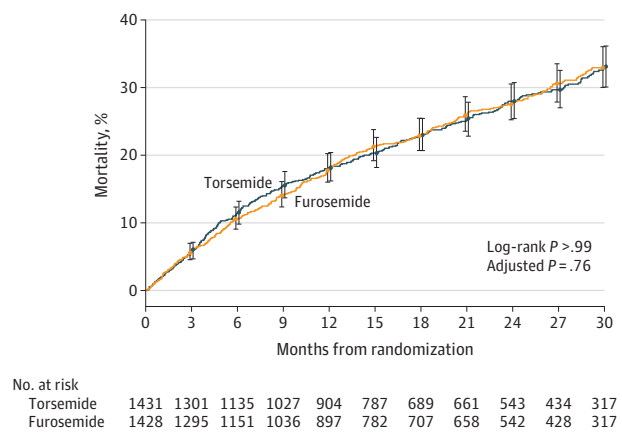
Adherence to Trial Medication

Of the patients with known prescription status at discharge from the index hospitalization (2755/2859, 96.4%), 2491 participants (90.4%) were receiving the assigned loop diuretic. At hospital discharge, we observed 7.0% crossover from toremide to furosemide and 3.8% crossover from furosemide to toremide (5.4% overall). In terms of loop diuretic discontinuation, 2.8% of patients were not discharged taking any loop diuretic. At 30 days and 6 months, 7.0% and 9.5% were not taking any loop diuretic, respectively. eTable 6 in Supplement 3 provides details of follow-up loop diuretic status.

Loop Diuretic Dosing

At index hospitalization discharge, the mean (SD) loop diuretic dose in furosemide equivalents (using a 2:1 conversion for furosemide to toremide) was 79.3 (63.3) mg and was similar in both groups for those individuals prescribed the assigned loop diuretic (79.1 [56.4] mg of furosemide vs 79.5 [69.8] mg of toremide). At 1 month (data available in 2047 participants, excluding deaths and unknown diuretic status), in those prescribed the assigned loop diuretic, the mean (SD) dose was 73.1 (63.4) mg, with a lower dose in the furosemide group than the toremide group (68.4 [50.2] mg vs 77.8 [74.5] mg).

Figure 2. Primary Outcome of All-Cause Mortality



The cumulative incidence of the primary outcome in the 2 groups is shown. The whiskers represent 95% CIs at the months specified. Variables for the adjusted P value are listed in Table 2.

Discussion

Among patients discharged after hospitalization for heart failure, toremide did not result in a significant difference in all-cause mortality compared with furosemide. However, interpretation of these findings is limited by loss to follow-up and participant crossover and nonadherence.

While prior mechanistic studies, observational analyses, and meta-analyses suggested advantages with toremide, this study did not demonstrate a treatment benefit compared with furosemide. There was no evidence that toremide's favorable bioavailability or purported antifibrotic effects translated into improved outcomes for patients recently hospitalized with heart failure. The results were consistent across the various end points and subgroups including those of different demographic profiles (eg, age, sex, race and ethnicity) and ejection fraction phenotypes and in those whose index hospitalization was with newly diagnosed heart failure as compared with worsening chronic heart failure. However, the non-specific all-cause outcomes may have been too imprecise for measuring subtle differences between the treatment groups. Furthermore, given that approximately 30% of the participants had newly diagnosed heart failure, postbaseline changes in guideline-directed medical therapy (eg, addition of ARNI, β -blocker, SGLT2 inhibitor, and mineralocorticoid receptor antagonist) may have affected clinical outcomes. In particular, the uptake of newer therapies, such as ARNI and SGLT2 inhibitor, over the course of the trial warrants consideration because these not only reduce clinical events, but potentially also diuretic requirements. Whether other patient populations, such as those diagnosed in the outpatient setting and/or without prior hospitalization, have differential benefit with these therapies was not assessed.

Crossover and loop diuretic discontinuation may have diminished the ability to distinguish the hypothesized between-group difference. Loop diuretics were prescribed as part of the

Table 2. Primary and Secondary Outcomes

Variable	Torsemide (n = 1431)		Furosemide (n = 1428)		Risk reduction (95% CI) ^a	HR (95% CI) ^b	P value ^b
	No. (%)	Events per 100 patient-years	No. (%)	Events per 100 patient-year			
Primary outcome							
All-cause mortality	373 (26.1)	17.0	374 (26.2)	17.0	0.12 (-2.85 to 3.14)	1.02 (0.89 to 1.18)	.76
Secondary outcomes							
All-cause mortality or all-cause hospitalization (over 12 mo)	677 (47.3)	99.2	704 (49.3)	107.6	1.99 (-1.79 to 5.56)	0.92 (0.83 to 1.02)	
Total hospitalizations (over 12 mo)	940	106.3	987	111.9		RR, 0.94 (0.84 to 1.07)	
All-cause mortality or all-cause hospitalization (over 30 d)	149 (10.4)	147.2	157 (11.0)	157.5	0.58 (-1.80 to 2.75)	0.94 (0.75 to 1.18)	

Abbreviations: HR, hazard ratio; RR, rate ratio.

^a Risk reduction = furosemide % minus torsemide %, and the 95% CI based on 2.5th and 97.5th percentiles of the risk reductions from 10 000 bootstrap samples of each treatment group.

^b HRs, 95% CIs, and P values are based on a Cox proportional hazards regression

model including the assigned treatment (torsemide vs furosemide as the reference group), as well as age, sex, baseline ejection fraction (<40%, 41%-49%, >50%, unknown), and loop diuretic treatment prior to index hospital admission as covariates.

routine strategy of care in the trial with mechanisms in place to support adherence to the randomized therapy, yet crossover was observed. Higher crossover in the torsemide group than furosemide group may relate to reversion to prior loop diuretic, differences in cost between the agents, patient or clinician preference, or perceived adverse effects. Despite the inclusion criterion of anticipated need for long-term loop diuretic therapy, diuretic discontinuation during follow-up was higher than expected for a hospitalized population. The as-treated analyses at discharge and 30 days supported the primary trial results. However, given challenges contacting patients at follow-up, there was missing data for follow-up diuretic status, which limits the interpretation.

There were significant differences in the loop diuretic dosing with regard to furosemide equivalents during follow-up. While dosing was similar at index hospital discharge, in patients with dosing data available at 1 month, dose was 10% to 15% greater in the torsemide than furosemide group (based on a 2:1 conversion) for participants continuing to take the randomized therapy. Given the uncertainty regarding the correct dose conversion, the protocol allowed flexibility with a 2:1 to 4:1 furosemide to torsemide conversion. If the true conversion is closer to 4:1, it may be that dosing was higher in the torsemide group. Future work will explore different dose conversions and time-varying analyses to better understand the implications of differences in dose.

The broad eligibility criteria, site selection, and streamlined study protocol embedded within routine care supported inclusion of diverse participants. By having centralized follow-up without site-specific visits, this supported inclusion of patients who historically have been less well represented in trials. In the study, 36.9% of trial participants were women and 33.9% were Black. Prior trials of patients with heart failure typically recruited less than 30% women¹⁹ and among trials reporting race, persons who were Black, Indigenous, or of racial or ethnic minority groups represented only 18.7% of study populations.²⁰ The pragmatic elements lowered traditional barriers for patient and site participation in clinical trials

and supported robust enrollment rates (even during the COVID-19 pandemic) with results that are generalizable to practice. A mean recruitment rate of more than 2 patients per site per month prior to the pandemic and more than 1 patient per site per month during the pandemic (following initial lockdown) highlight advantages of trials incorporating centralized follow-up mechanisms to reduce the burden on enrolling sites and patients.

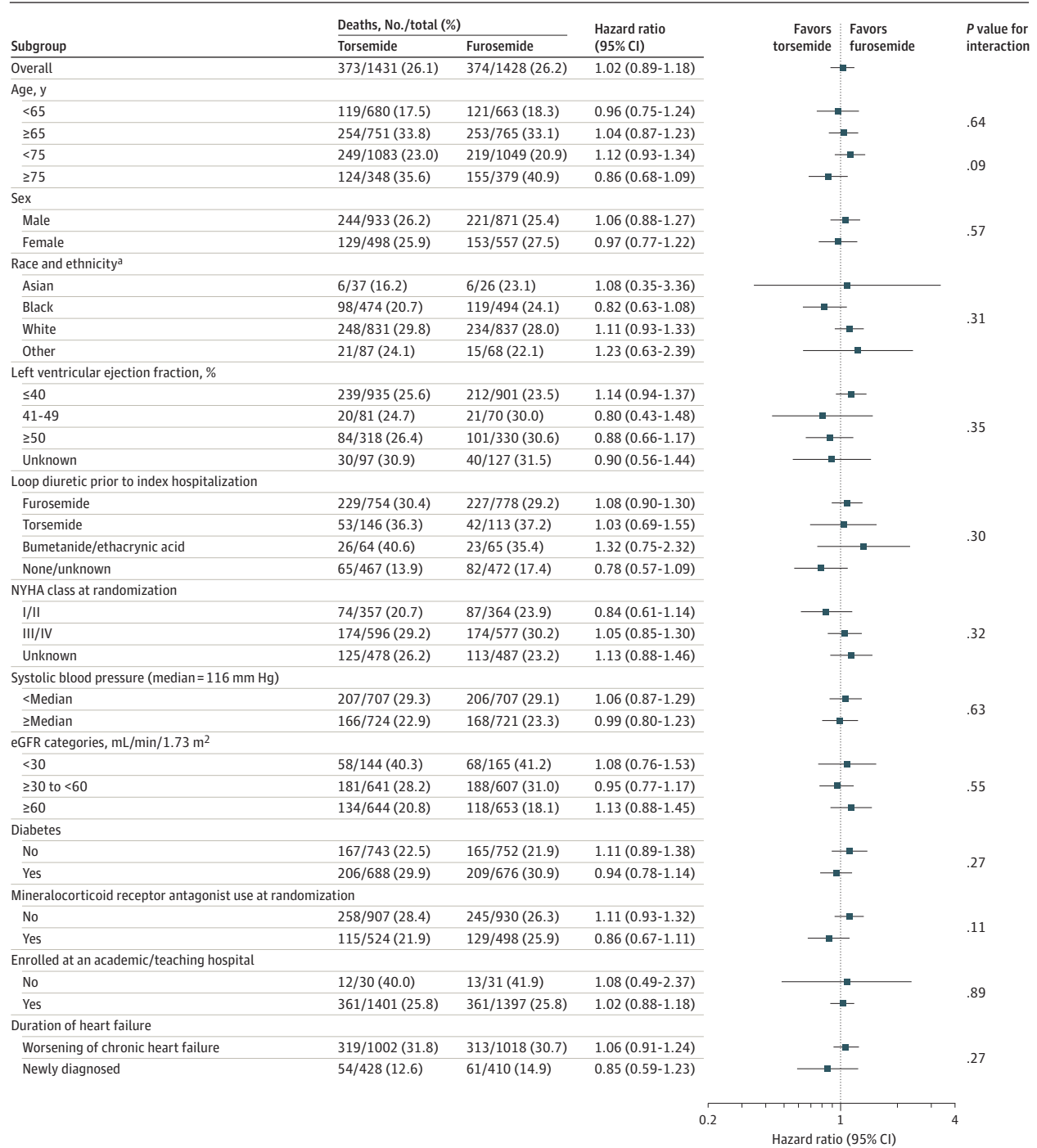
The event-driven trial was initially projected to target the recruitment of 6000 participants. However, the broad eligibility criteria and in-hospital recruitment preceding the vulnerable period post discharge²¹ supported a higher-than-anticipated event rate. The trial reached the target event count of 721 death events (747 observed deaths) with a sample size approximately half that initially planned. The event rate of 17.0 per 100 patient-years was similar to that observed in the recent clinical outcome trial of vericiguat following a recent worsening heart failure event (placebo all-cause mortality rate of 16.9 per 100 patient-years in the VICTORIA trial).²² This study enrolled a high-risk phenotype with a median N-terminal pro-brain natriuretic peptide level of 3913 pg/mL as compared with 4812 pg/mL in the PIONEER-HF trial,²³ 2816 pg/mL in the VICTORIA trial,²² and 1437 pg/mL in the DAPA-HF trial.²⁴ The high baseline N-terminal pro-brain natriuretic peptide in this study is notable given that the trial included patients with preserved ejection fraction where natriuretic peptide levels are comparatively lower than in reduced ejection fraction.

Limitations

Several important limitations should be acknowledged. First, while we achieved the target event count, the sample size was approximately half that originally planned. Subgroup analyses were therefore limited by the more modest patient numbers. Patient withdrawals were higher than in some prior heart failure trials, likely related in part to the reduced intensity of site contact.

Second, the all-cause outcomes may have been too imprecise for measuring subtle differences between the treatment

Figure 3. Primary Outcome in Prespecified Subgroups



Results of the primary outcome of the trial—all-cause mortality—are shown according to subgroups that were prespecified in the protocol. eGFR indicates estimated glomerular filtration rate and NYHA class, New York

Heart Association symptom class at time of randomization. ^a American Indian, Native Hawaiian, and "multiple" categories were not included due to very small sample sizes.

groups in heart failure-specific outcomes. The end point classification differed from traditional clinical outcome adjudication with cause-specific end points given the pragmatic nature of the trial. The treatment effect assumed in trial planning was informed by available meta-analyses, which may have not

fully incorporated the beneficial effects of guideline-directed medical therapy (particularly for reduced ejection fraction heart failure).

Third, while clinical events were systematically evaluated, the pragmatic design did not allow for assessment of

subtle relative benefits (or harms) of these generically available therapies such as worsening kidney function, electrolyte abnormalities, or nonhospitalization events (eg, emergency department visits, outpatient intravenous diuretics, thiazide use).

Fourth, loop diuretic discontinuation and crossovers occurred during follow-up and are informative in this comparative strategy study, yet they may bias toward neutral results. Given the open-label design, it is plausible that patient or clinician bias about differential benefits of the loop diuretics may have led to switching over time.

Fifth, loop diuretic dose was left to clinician discretion, which may have influenced results. Future work will characterize how nonadherence and dose titration may have af-

ected these findings including via the evaluation of varying definitions of “as-treated” and incorporation of time-varying factors.

Sixth, the recruitment of patients with heart failure with preserved ejection fraction and individuals of Hispanic ethnicity was lower than anticipated.

Conclusions

Among patients discharged after hospitalization for heart failure, torsemide compared with furosemide did not result in a significant difference in all-cause mortality over 12 months. However, interpretation of these findings is limited by loss to follow-up and participant crossover and nonadherence.

ARTICLE INFORMATION

Accepted for Publication: December 9, 2022.

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Obtained funding: Mentz, Testani, Velazquez.

Administrative, technical, or material support: Mentz, Eisenstein, Morgan, Testani, Harrington, Ketema, Desvigne-Nickens, Velazquez.

Supervision: Mentz, Harrington, Sachdev, Velazquez.

Conflict of Interest Disclosures: Dr Mentz reported receiving grants from American Regent and Novartis; personal fees from AstraZeneca, Boehringer Ingelheim/Eli Lilly, Cytokinetics, Bayer, Merck, and Pharmacosmos; and research support from Abbott, Amgen, Bayer, Boston Scientific, Fast BioMedical, Gilead, Innolife, Medtronic, Relypsa, Respicardia, Roche, Sanofi, Vifor, Windtree Therapeutics, and Zoll outside the submitted work. Dr Anstrom reported receiving grants from Merck,

Bayer, the National Institutes of Health, and the Patient-Centered Outcomes Research Institute during the conduct of the study. Dr Eisenstein reported receiving grants from the National Heart, Lung, and Blood Institute (NHLBI) during the conduct of the study. Dr Sapp reported receiving grants from Duke Clinical Research Institute during the conduct of the study. Dr Greene reported receiving grants from the NHLBI and American Heart Association during the conduct of the study; personal fees from Amgen, AstraZeneca, Bayer, Bristol Myers Squibb, Boehringer Ingelheim/Eli Lilly, Criterion Therapeutics, Cytokinetics, Merck, PharmaIN, Roche Diagnostics, Tricog Health, Sanofi, Urovant Pharmaceuticals, and CSL Vifor and nonfinancial support from AstraZeneca, Bristol Myers Squibb, Cytokinetics, Merck, Novartis, Sanofi, scPharmaceuticals, and Pfizer outside the submitted work. Dr Testani reported receiving grants from the National Institutes of Health during the conduct of the study; personal fees from 3ive Labs, AstraZeneca, Novartis, Cardionomic, MagentaMed, Reprieve Inc, Sanofi, Sequana Medical, Merck, Windtree Therapeutics, Lexicon Pharmaceuticals, Precardia, Relypsa, Regeneron, BD, Edwards Lifesciences, Eli Lilly, Bayer, Boehringer Ingelheim, Bristol Myers Squibb, and W. L. Gore, and grants from FIRE1, Sequana Medical, Otsuka, Abbott, Merck, Lexicon Pharmaceuticals, 3ive Labs, Boehringer Ingelheim, Bristol Myers Squibb, and Sanofi outside the submitted work. In addition, Dr Testani had a patent for treatment of diuretic resistance with royalties paid from Yale and Corvidia Therapeutics, a patent for formulations and methods for direct sodium removal in patients having severe renal dysfunction issued from Reprieve Inc, and a patent for formulations and methods for direct sodium removal in patients having severe renal dysfunction issued from Sequana Medical. Dr Pitt reported receiving personal fees from Bayer, scpharmaceuticals, Sqinnovation, Boehringer Ingelheim, Vifor, AstraZeneca, G3 Pharmaceuticals, Sarfez, Cereno Scientific, Phasebio, KPB Biosciences, Protonintel, Brainstorm Medical, and Merck outside the submitted work. In addition, Dr Pitt had a patent for US 9931413 issued, site specific delivery of eplerenone to the myocardium, and a patent for US 63/045,783 pending, histone modulating agents for the prevention and protection from organ injury. Dr Velazquez reported receiving grants from the NHLBI to Yale University as principal investigator of the clinical coordinating center during the conduct of the

study; and personal fees from Novartis and consultant fees from Abiomed to Yale University and from Baim Institute to Yale University outside the submitted work. No other disclosures were reported.

Funding/Support: TRANSFORM-HF is supported through cooperative agreements from the NHLBI (U01-HL125478 and U01-HL125511) and ancillary study grants from the NHLBI (R01HL148354-04 and R01HL154768-02).

Role of the Funder/Sponsor: The NHLBI had no role in the design of the study but NHLBI staff did participate in the study conduct, collection, management, analysis, and interpretation of the data; preparation, review, or approval of the manuscript; and decision to submit the manuscript for publication as members of the executive committee (Drs Sachdev, Kim, and Desvigne-Nickens). The sponsor did not have the right to veto publication or to control the decision regarding to which journal the paper was submitted.

Group Information: A complete list of the TRANSFORM-HF Trial Investigators is provided in [Supplement 4](#).

Disclaimer: The views expressed in this article are those of the authors and do not necessarily represent the views of the NHLBI, the National Institutes of Health, or the US Department of Health and Human Services.

Data Sharing Statement: See [Supplement 5](#).

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